

Privacy in Statistics and Machine Learning **Spring 2025**
In-class Exercises for Lecture 1 (Intro and Randomized Response)
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Problems with marked with an asterisk () are more challenging or open-ended.*

- Suppose X_1, \dots, X_n are independent random variables, each with mean $\mathbb{E}(X_i) = \mu$ and standard deviation $\sigma = \sqrt{\text{Var}(X_i)}$ (for all i).

What are the expectation and variance of the average $\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$?

- $\mathbb{E}(\bar{X}) = \mu n$ and $\sqrt{\text{Var}(\bar{X})} = n\sigma$
- $\mathbb{E}(\bar{X}) = \mu$ and $\sqrt{\text{Var}(\bar{X})} = \sigma$
- $\mathbb{E}(\bar{X}) = \mu$ and $\sqrt{\text{Var}(\bar{X})} = \sigma/\sqrt{n}$
- $\mathbb{E}(\bar{X}) = \mu$ and $\sqrt{\text{Var}(\bar{X})} = \sigma/n$
- $\mathbb{E}(\bar{X}) = \mu/n$ and $\sqrt{\text{Var}(\bar{X})} = \sigma/n$

- Recall the randomized response mechanism discussed in class. For each input bit x_i , it generates

$$Y_i = \begin{cases} x_i & \text{w.p. } 2/3, \\ 1 - x_i & \text{w.p. } 1/3. \end{cases}$$

Give a procedure that, given the outputs Y_1, \dots, Y_n from randomized response on input x_1, \dots, x_n , returns an estimate A such that

$$\sqrt{\mathbb{E} \left(\left(A - \sum_{i=1}^n x_i \right)^2 \right)} = O(\sqrt{n})$$

Hint: Find a function f that rescales the Y_i so that $\mathbb{E}(f(Y_i)) = x_i$.

- (*) Suppose now that for each respondent we have some “public” information (that is, known to the analyst), together with the private bit x_i . We might be interested in solving some task that involves both the public and private information, such as finding a model to predict x_i given the public features.

Consider the following super-simplified version of this: suppose the public information for person i is a real number $a_i \in \mathbb{R}$. Given the Y_i output by randomized response, how can we get an unbiased estimate of $\sum_i a_i x_i$? What is its variance (as a function of the list of a_i 's)?

4. (*) Consider the second randomized response mechanism described in class, in which

$$Y_i = \begin{cases} x_i & \text{w.p. } \frac{e^\epsilon}{e^\epsilon + 1}, \\ 1 - x_i & \text{w.p. } \frac{1}{e^\epsilon + 1}. \end{cases}$$

Give a procedure that, given the outputs Y_1, \dots, Y_n from randomized response on input x_1, \dots, x_n , returns an estimate A such that $\sqrt{\mathbb{E} \left((A - \sum_{i=1}^n x_i)^2 \right)} = \frac{e^{\epsilon/2}}{e^\epsilon - 1} \sqrt{n}$.

Hint: Find a function f that rescales the Y_i so that $\mathbb{E}(f(Y_i)) = x_i$.

Reminders on sums of random variables A good reference on the probability material needed for this class is the book of Mitzenmacher and Upfal [MU17]. We include here a few reminders that will be useful in today's lecture.

- Expectations are linear: If X, Y are random variables (it does *not* matter if they are independent), then for any constants $a, b \in \mathbb{R}$, we have

$$\mathbb{E}(aX + bY) = a\mathbb{E}(X) + b\mathbb{E}(Y).$$

By induction, linearity extends to finite sums (so $\mathbb{E}(a_1X_1 + \dots + a_kX_k) = a_1\mathbb{E}(X_1) + \dots + a_k\mathbb{E}(X_k)$).

- Variances add when random variables are independent: For any *independent* random variables X, Y , and for any constants $a, b \in \mathbb{R}$, we have

$$\text{Var}(aX + bY) = a^2\text{Var}(X) + b^2\text{Var}(Y).$$

Again, by induction, if X_1, \dots, X_k are independent, then $\text{Var}\left(\sum_{i=1}^k a_i X_i\right) = \sum_{i=1}^k a_i^2 \text{Var}(X_i)$. Note that variances do not necessarily add for *dependent* random variables. For example, if $Y = -X$, what is the variance of $X + Y$?

- Chebyshev's inequality: For any random variable X with finite mean and variance, for every $t > 0$, we have

$$\mathbb{P}\left(|X - \mathbb{E}(X)| \geq t\sqrt{\text{Var}(X)}\right) \leq 1/t^2.$$

- "Chernoff bounds" are a family of concentration inequalities for sums of independent random variables. A useful example is the following:

Lemma 0.1. Let X_1, \dots, X_n be i.i.d. random variables taking values in $[0, 1]$. Let X denote their sum and let $\mu = \mathbb{E}(X_i)$ (so that $\mathbb{E}(X) = \mu n$). Then,

- For every $\delta \geq 0$, $\mathbb{P}(X > (1 + \delta)\mu n) \leq e^{-\delta^2 \mu n/3}$
- For every $\delta \in [0, 1]$, $\mathbb{P}(X < (1 - \delta)\mu n) \leq e^{-\delta^2 \mu n/2}$.

In particular, for every $t > 0$, the probability that $|X - \mu n| \geq t\sqrt{n}$ is at most $2 \exp(-t^2/3)$.

References

- [MU17] Michael Mitzenmacher and Eli Upfal. *Probability and Computing: Randomization and Probabilistic Techniques in Algorithms and Data Analysis*. Cambridge University Press, 2nd edition, 2017.